Observed structure of addresses in IP traffic



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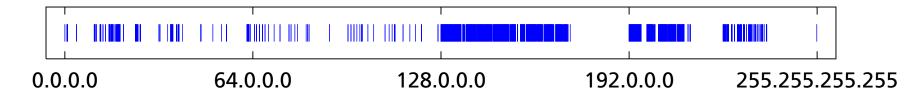
Thanks to David Donoho and Dick Karp

Problem



 How can we model the set of destination IP addresses visible on some link? (And does it matter?)

Example from a 4-hour trace at a university access link:



In particular, can we model how the addresses aggregate?

We call this address structure.

• Applications might include average-case route lookup, analysis of aggregate-based congestion control, realistic sets of addresses for simulations, . . .

Results



- Address structure dominates the characteristics of medium-scale prefix aggregates, such as /16s.
- The medium-scale aggregation behavior of real addresses is well modeled by a multifractal Cantor set construction with two parameters.

The model captures both fractal metrics and metrics we developed for address structures.

- Address structure can serve as a site "fingerprint".
 - Structural metrics differ between sites.
 - At a given site, these metrics are stable over short time scales.
 - New communication dynamics, such as worm propagation, show up in the metrics.

Outline



- Terminology
- Address structure and aggregate packet counts
- Model
- Metrics
- Fingerprints

Terminology



- Active address: an IP address visible in the trace as a destination
- N: the number of active addresses in a trace $N \le 2^{32}$ by definition; $N \ll 2^{32}$ for all our traces
- p-aggregate: a set of addresses that share the same p-bit address prefix (0 $\leq p \leq$ 32)

Also called a /p

1.0.0.0 and 1.99.130.14 are in the same /8, but different /10s

• Active p-aggregate: a /p containing at least one active address

Traces



Name	Description	ΔT	# pkts	N	
U1	large university access link	\sim 4 h	62M	69,196	
U2	large university access link	\sim 1 h	101M	144,244	
A1	ISP	\sim 0.6 h	34M	82,678	
A2	ISP	1 h	29M	154,921	
R1	link from regional ISP	1 h	1.5M	168,318	§
R2	link from regional ISP	2 h	1M	110,783	§
W1	large Web site access link	\sim 2 h	5M	124,454	

• Collected between 1998 and 2001

Most anonymized while preserving prefix and class relationships § means sampled (1 in 256)

Does address structure matter?



Assume that aggregate packet counts matter.

Accounting, fairness, congestion control . . .

What factors affect aggregate packet counts?

Packet counts per address: probably a heavy-tailed distribution

Addresses per aggregate = address structure

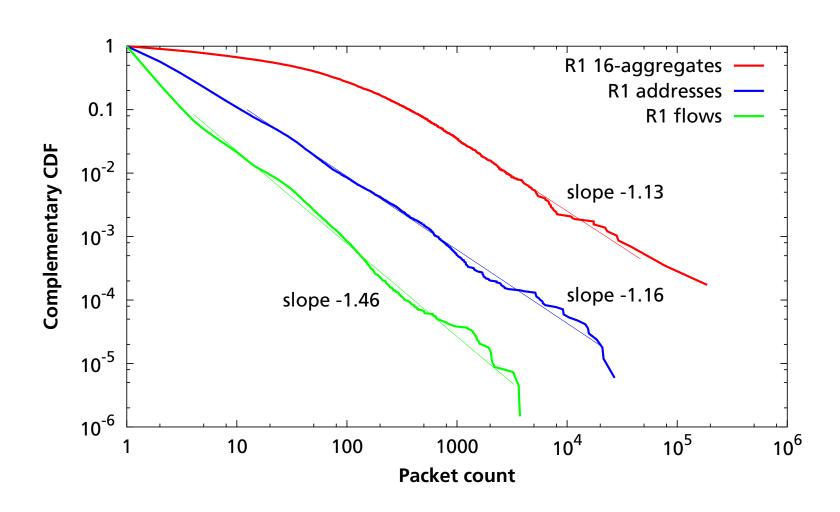
Correlation

 Analyze the contributions of these factors to an observed packet count distribution

Medium scales are most interesting (/16s and thereabouts)

R1 packet count distributions





Semi-experiments



- Manipulate the data, destroying one factor at a time; see which factors impact aggregate packet counts
- "Random counts": destroy per-address packet counts

Replace the (heavy-tailed) per-address packet count distribution with a uniform distribution over [0, 17.54]

• "Random addresses": destroy address structure

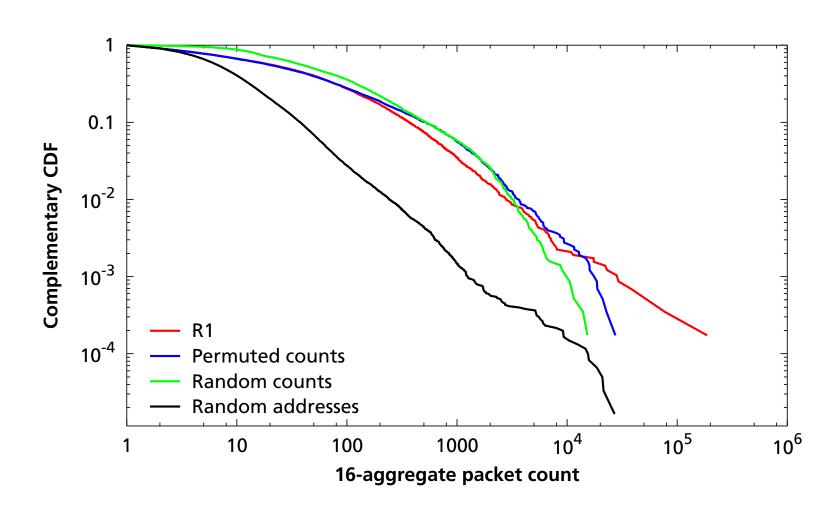
Replace address structure with a uniform random distribution over the entire IP address space

"Permuted counts": destroy correlation

Permute per-address packet counts among the active addresses

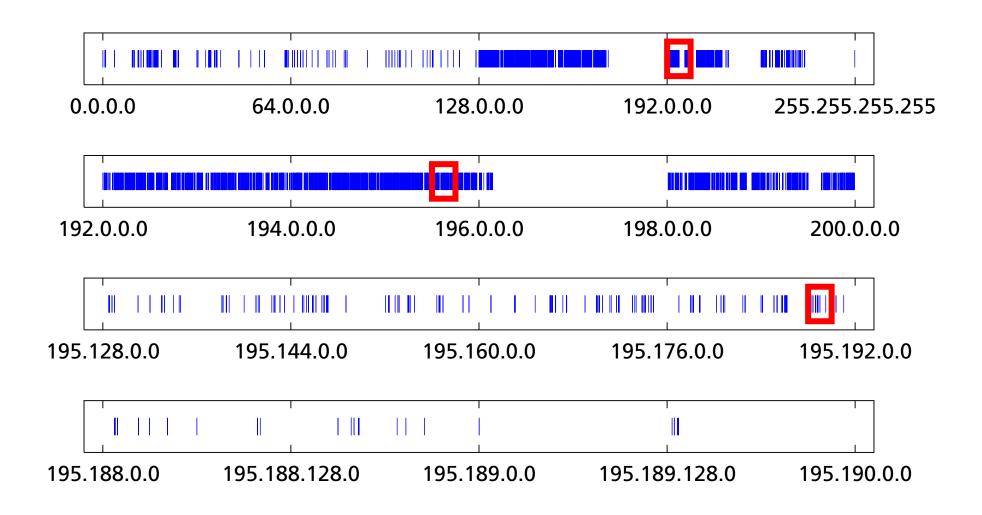
Address structure matters most





Tour of U1's address structure





Self-similarity?



- Interesting structure all the way down
 - Visually "self-similar" characteristics
- Might address structure be usefully modeled by a fractal?
 - Treat an address structure as a subset of the unit interval
 - Fractal dimension $D \in [0, 1]$?

Fractal dimension for address structure



- Use lattice box-counting dimension
 - Corresponds nicely to prefix aggregation
- Let n_p equal the number of active /ps in a trace

$$n_{32} = N$$

$$n_p \leq n_{p+1} \leq 2n_p$$

each /p contains and is covered by 2 disjoint /(p + 1)s

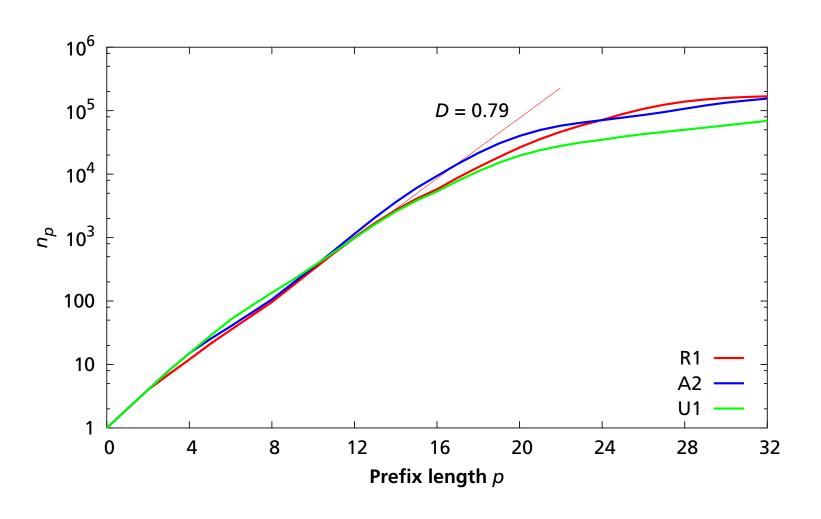
• Then
$$D = \lim_{p \to \infty} \frac{\log n_p}{p \log 2}$$

But $p \leq 32$ here, and expect sampling effects for high p

Examine medium p to see if the limit exists

$\log n_p$ is linearly related to p at medium scales





Multifractality



Monofractal may not be sufficient

Same scaling behavior everywhere

Not what we saw in the tour

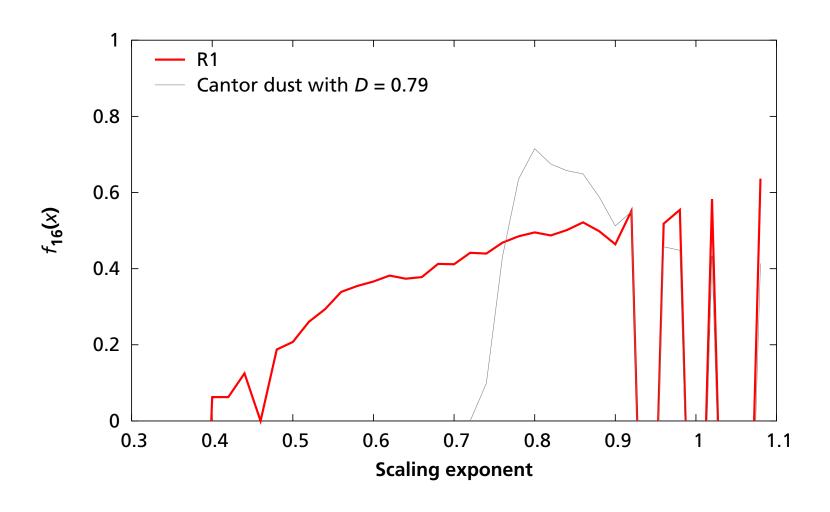
• Examine the *multifractal spectrum* to test for multifractality (different local scaling behavior)

Binned approximation (Histogram Method)

If multifractal, spectrum will cover a wide range of scaling exponents

Address structure is multifractal at /16





Multifractal model



- Make a multifractal Cantor measure matching this spectrum
- Start with a Cantor dust with dimension DRepeatedly remove middle subinterval with proportion $h=1-2^{1-1/D}$
- Sample unequally from left and right subintervals

Distribute a unit of "mass" between subintervals; left gets m_0 , middle gets 0 (removed), right gets $m_2 = 1 - m_0$

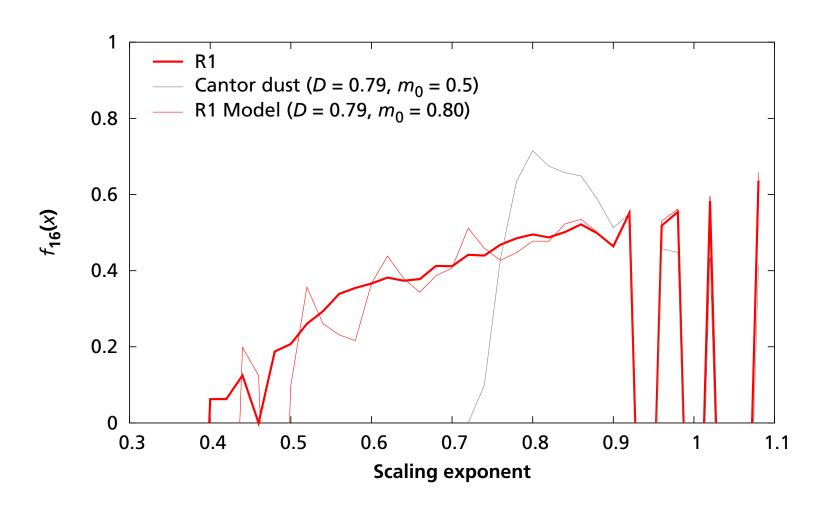
Produces a sequence of measures μ_k that weakly converge to μ

Sample an address with probability equal to its measure

Result: different local scaling behavior

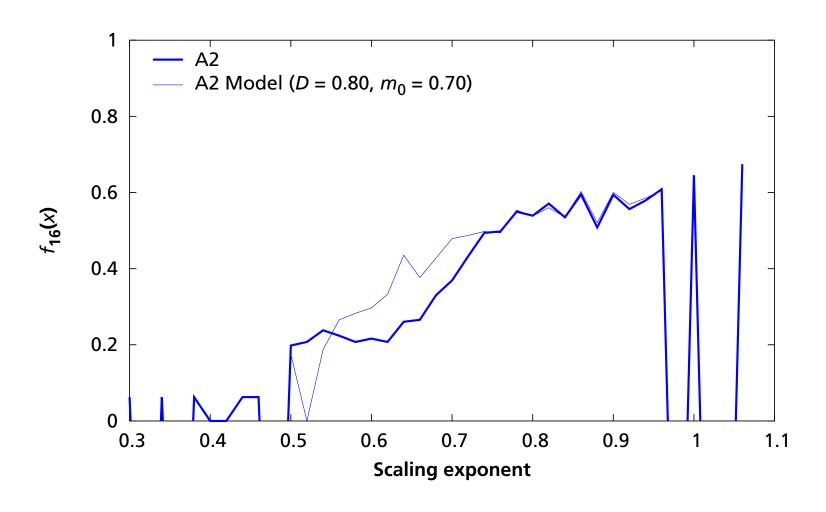
The model fits well





The model fits well





Why multifractal?



Perhaps it's due to a cascade

Recursive subdivision plus a rule for distributing mass

- For example, address allocation
 - Pure speculation!
 - ICANN allocates short prefixes to providers
 - Providers allocate longer prefixes to their customers
 - All parties might allocate basically from left to right

Does the multifractal spectrum matter?



Certainly the model doesn't look like real data:



How do we know whether we've captured relevant properties?

- Develop application metrics for address structures
 - Contrast metrics among traces
 - Compare with model

Active aggregate counts: n_p and γ_p



- n_p again equals the number of active /ps in a trace
- n_p measures how densely addresses are packed If $N=2^{16}$ and $n_{16}=1$, addresses are closely packed

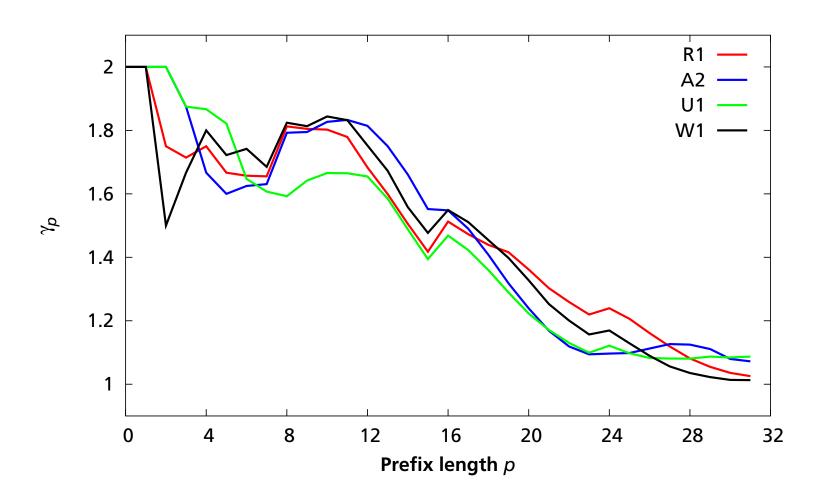
If $N=2^{16}$ and $n_{16}=2^{16}$, addresses are well spread out

Useful for algorithms keeping track of aggregates—shows how many aggregates there tend to be

• $\gamma_p = n_{p+1}/n_p$ more convenient for graphs $N = \prod_{1$

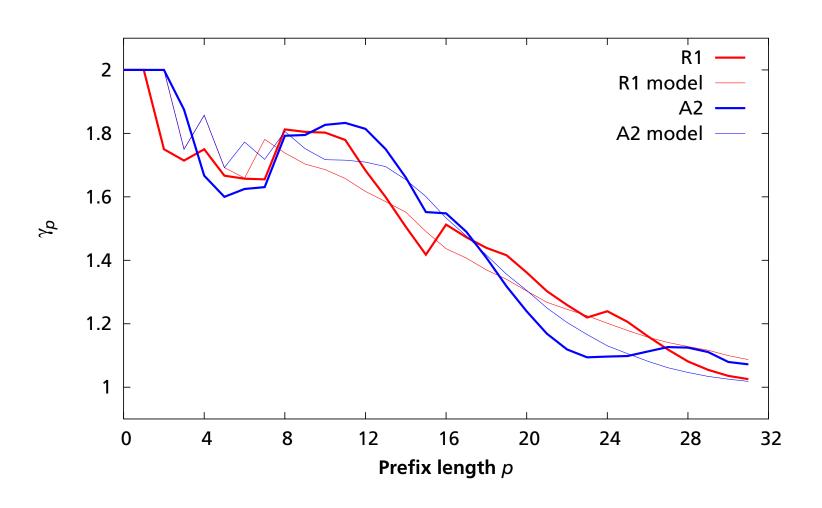






Models' γ_p





Discriminating prefixes



• The **discriminating prefix** of an active address, *a*, is the prefix length of the *largest* aggregate that contains *only one* active address, namely *a*.

Example with 4-bit addresses:

gth	4	4 4	3	2	4 4	
<u>le</u> n	3					
fix	1					
Pre	0					

• Measures address separation

If many addresses have d.p. < 20, say, then addresses are well separated

How depopulated do aggregates become?

Discriminating prefixes: π_p



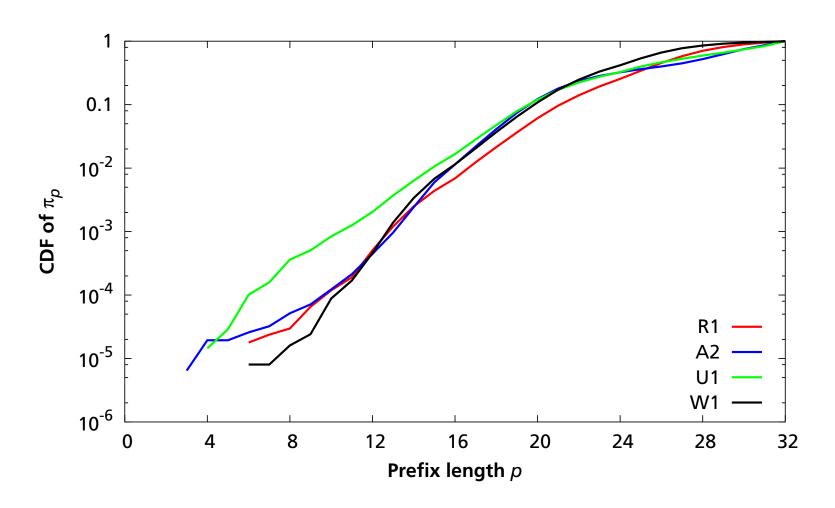
• Let π_p equal the number of addresses with d.p. p

$$\sum \pi_{p} = N$$

Turns discriminating prefixes into a metric

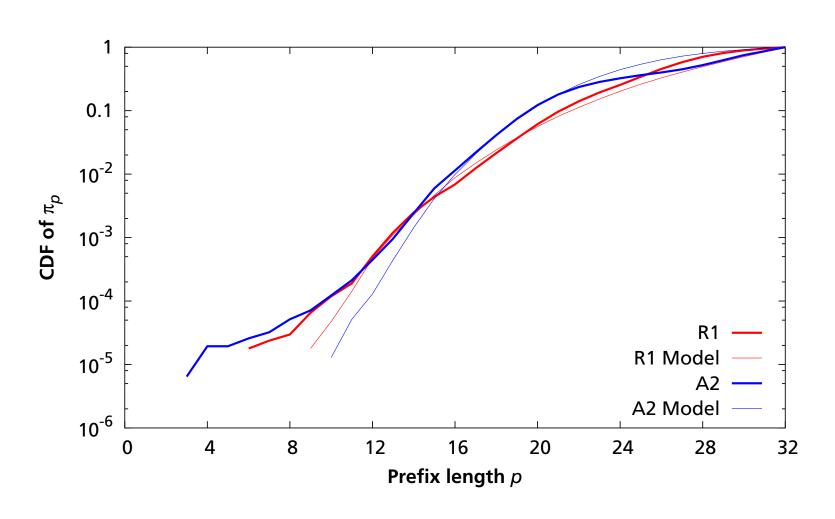
 π_{p}





Models' π_p





Aggregate population distribution

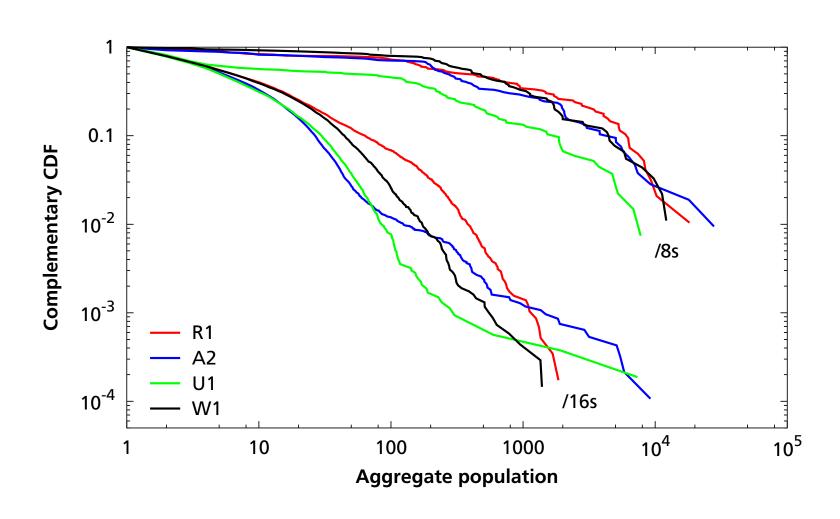


• Like aggregate packet count distribution, but count the number of active addresses per aggregate

Expect a wide range of variation, just as with the other metrics

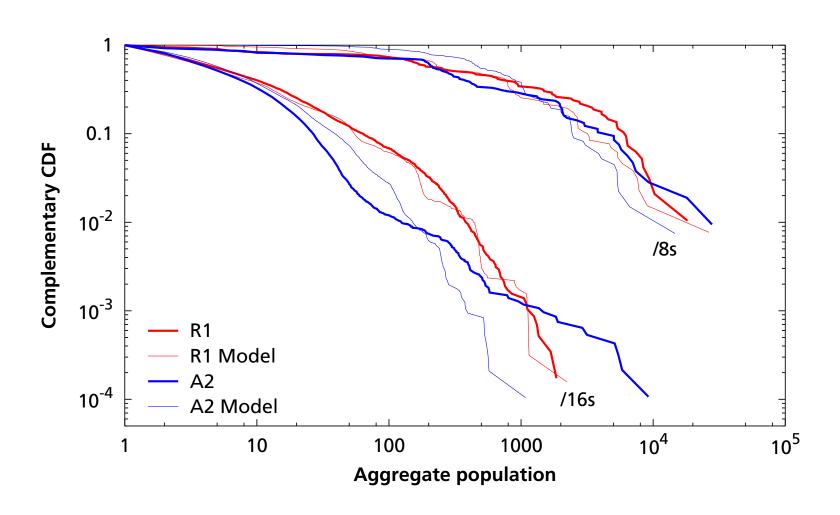
Aggregate population distribution





Models' aggregate population distribution





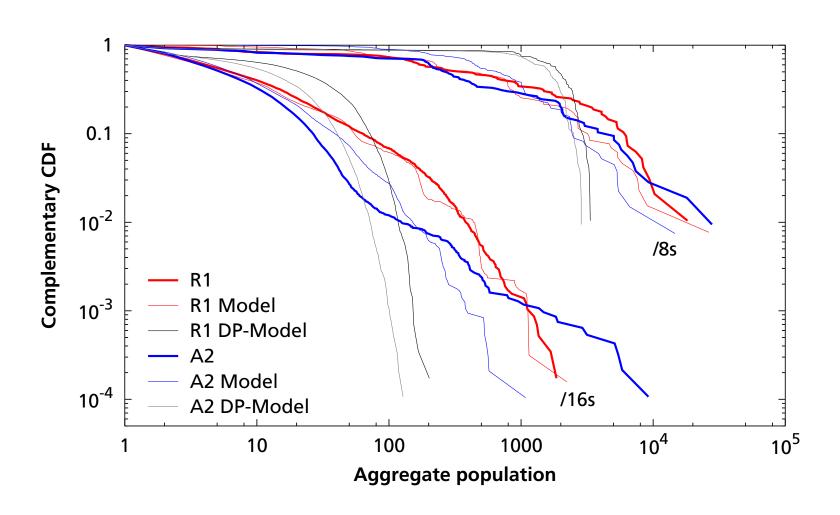
A tough metric



- The model for A2 doesn't match A2's aggregate populations
 - R1, W1 match well, A2, U1 do not
 - Significant aggregation in A2, U1 at long prefixes . . . ?
- Aggregate population distribution is difficult to match
- Consider random allocation constrained to match γ_p and π_p exactly
 - Heck, match "generalized discriminating prefixes"—d.p.s for
 - aggregates—as well
 - Call this the "Match-DP" model
 - How well does this do?

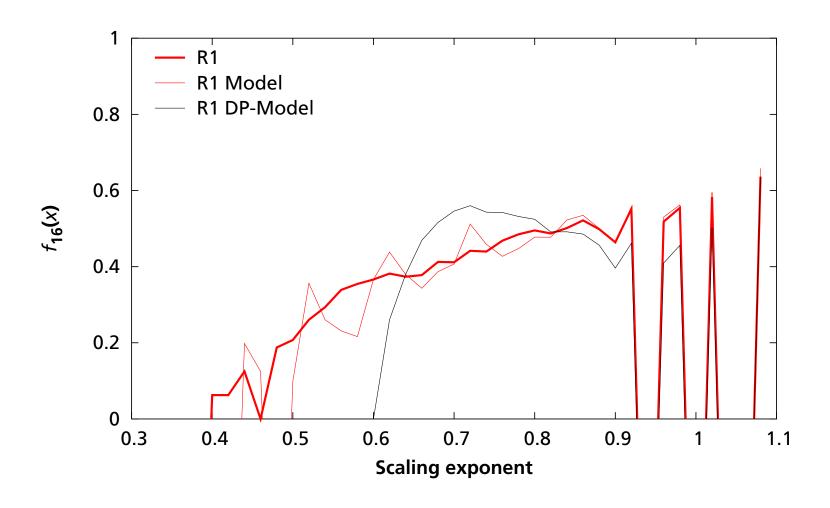
Match-DP fails aggregate population distribution





Another tough metric: The multifractal spectrum





Properties of γ_p : Sampling effects?



- Turn from the multifractal model to properties of our γ_p metric
- First: Is γ_p dominated by sampling effects?

N is effectively a sample size

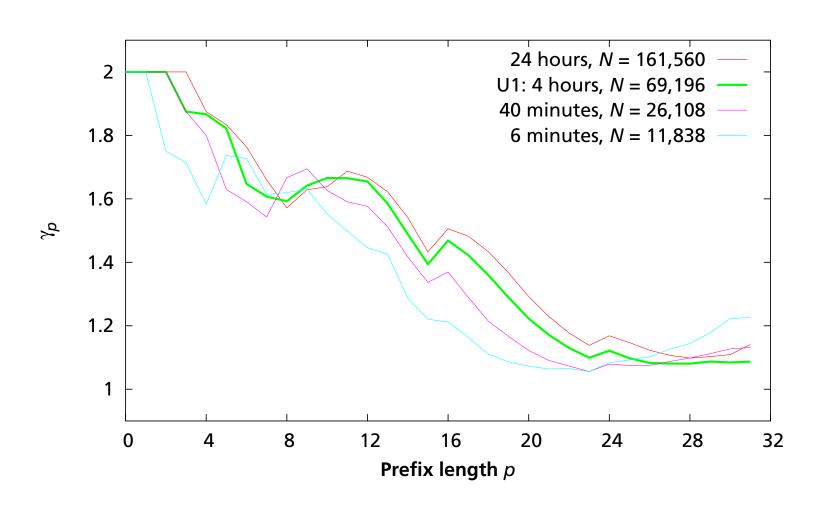
How does the shape of the γ_p curve depend on N?

• Plot γ_p for longer and shorter sections of trace U1

24 hours
$$\rightarrow$$
 6 minutes; **N** = 161,560 \rightarrow 11,838

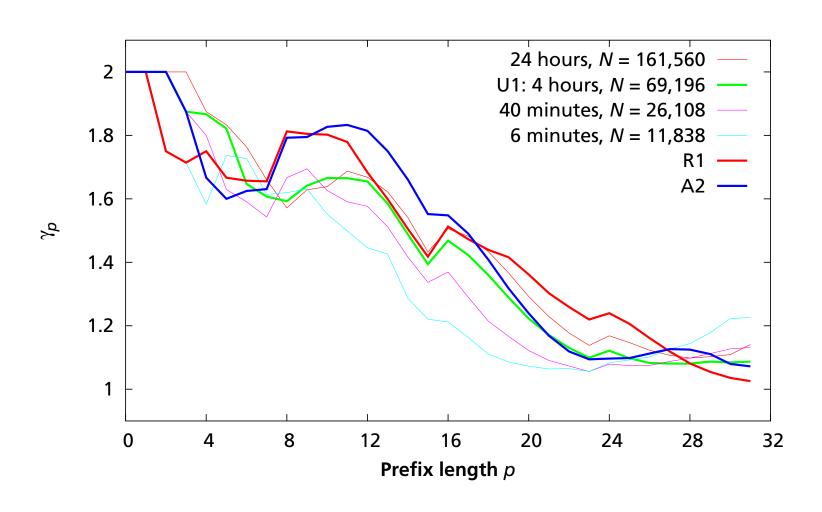
Shape of γ_p similar for wide range of sample sizes





Shape of γ_p similar for wide range of sample sizes





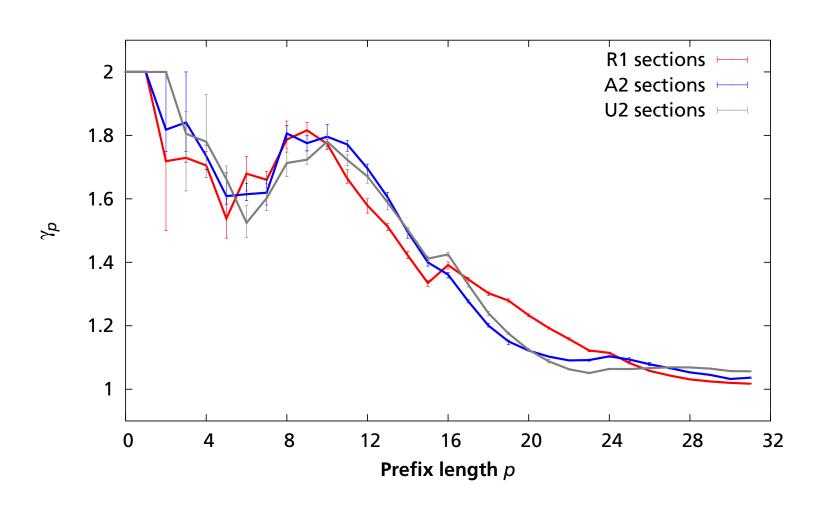
Short-term stability?



- Is γ_p stable over short time scales?
- Divide traces into short sections, each with N=32,768Plot maximum, minimum, and mean γ_p over all sections R1, A2, and U2; sections last about 6–7 minutes each

Shape of γ_p relatively stable over short time scales





New communication dynamics?



• How does γ_p change given a different communication pattern, such as worm propagation?

Expect worm propagation to significantly change the destination addresses visible at an access link, since every possible internal address will be contacted.

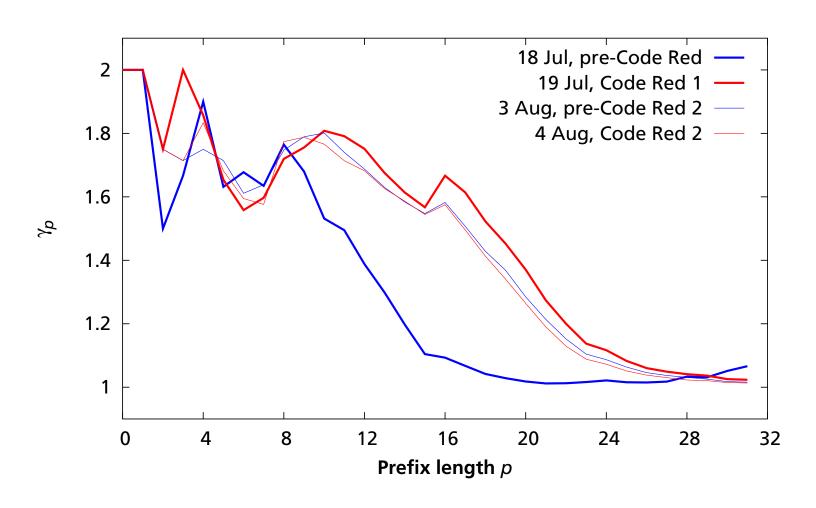
Not the best detection metric . . .

 Take a new data set, collected at a national laboratory, before and after Code Reds 1 and 2

Consider γ_{p} and aggregate population distribution

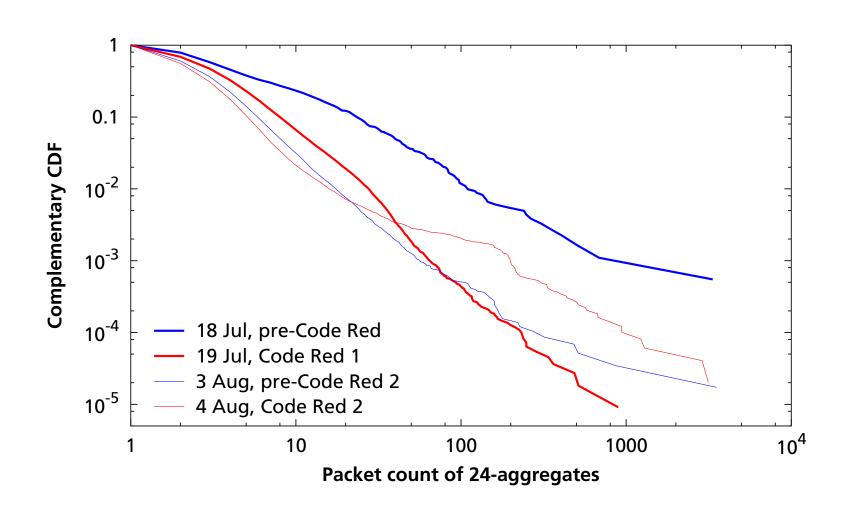
Shape of γ_p changes during worm propagation





Agg. packet counts change during worm propagation





Address stability



- Divide a trace into sections, each lasting *t* seconds.
- How many addresses in section 1 recur in section 2?
 - ... in sections 1, 2, and 3? and so forth Indicates how quickly address sets change
- Model: there are long-lived addresses and short-lived addresses

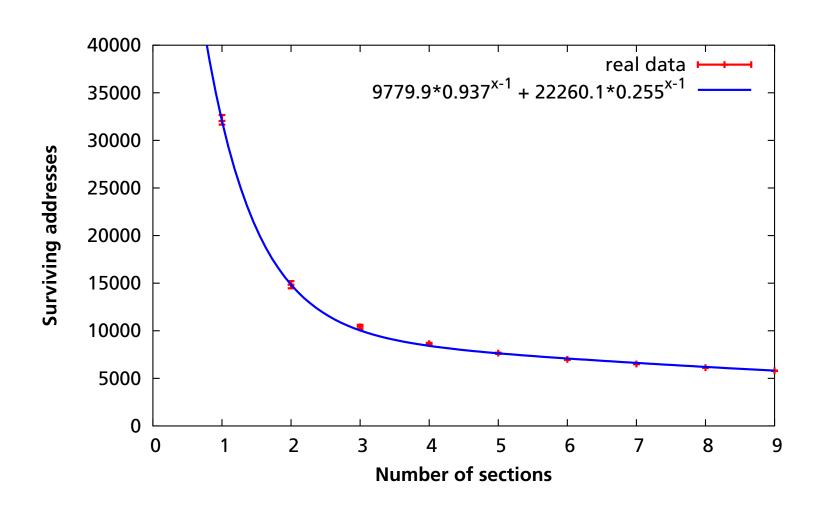
Every section contains n_S short-lived and n_L long-lived

Addresses survive into the next section with probabilities ho_S and ho_L (where $ho_L >
ho_S$)

How well does this model match?

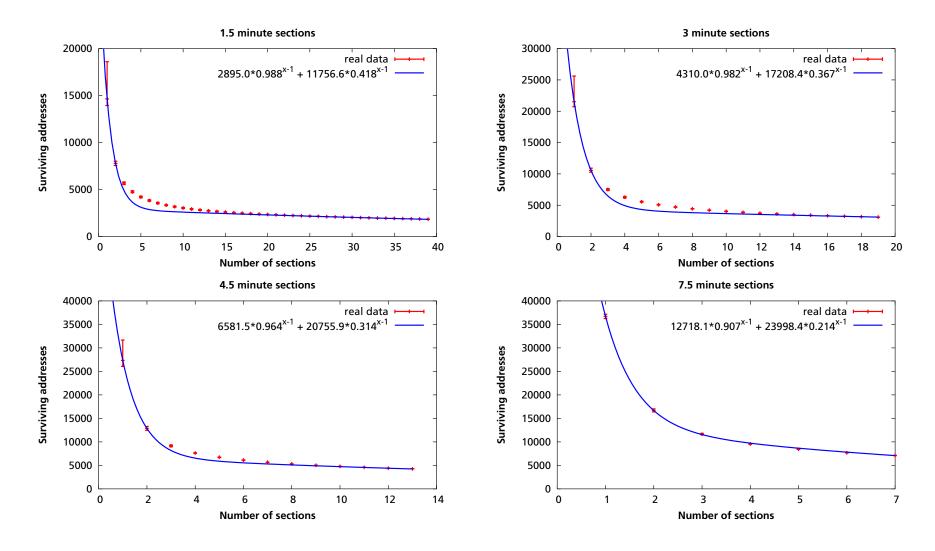
U2, 6-minute sections





Other time scales





Conclusions



- Demonstrated importance of address structure
- Real address structure well modeled by a two-parameter multifractal

Captures some aggregation behavior better than models built using metrics from real data

• Use of structural metrics as site fingerprints

Metrics differ between sites, are stable over short time scales

Future work





Analysis details



- Sections are numbered 1...k.
 n[A] is number of active addresses in intersection of sections A.
- n_L long-lived addresses per section, n_S short-lived addresses.
- p_L long-lived survival probability, p_S short-lived.
- $p_L \sim n[1...k]/n[1...k-1]$.
- $n_L = n[1...k]/p_L^k$.
- $n_S = n[1] n_L$.
- $p_S = (n[1, 2] n_L p_L)/n_S$.